RESEARCH ARTICLE

BMC Medicine



Plasma metabolites and risk of seven cancers: a two-sample Mendelian randomization study among European descendants



Yaxin Chen^{1,2}, Yufang Xie², Hang Ci¹, Zhengpei Cheng³, Yongjie Kuang⁴, Shuqing Li¹, Gang Wang⁵, Yawen Qi¹, Jun Tang¹, Dan Liu², Weimin Li^{1,2*} and Yaohua Yang^{3*}[®]

Abstract

Background While circulating metabolites have been increasingly linked to cancer risk, the causality underlying these associations remains largely uninterrogated.

Methods We conducted a comprehensive 2-sample Mendelian randomization (MR) study to evaluate the potential causal relationship between 913 plasma metabolites and the risk of seven cancers among European-ancestry individuals. Data on variant-metabolite associations were obtained from a genome-wide association study (GWAS) of plasma metabolites among 14,296 subjects. Data on variant-cancer associations were gathered from large-scale GWAS consortia for breast (N = 266,081), colorectal (N = 185,616), lung (N = 85,716), ovarian (N = 63,347), prostate (N = 140,306), renal cell (N = 31,190), and testicular germ cell (N = 28,135) cancers. MR analyses were performed with the inverse variance-weighted (IVW) method as the primary strategy to identify significant associations at Bonferroni-corrected P < 0.05 for each cancer type separately. Significant associations were subjected to additional scrutiny via weighted median MR, Egger regression, MR-Pleiotropy RESidual Sum and Outlier (MR-PRESSO), and reverse MR analyses. Replication analyses were performed using an independent dataset from a plasma metabolite GWAS including 8,129 participants of European ancestry.

Results We identified 94 significant associations, suggesting putative causal associations between 66 distinct plasma metabolites and the risk of seven cancers. Remarkably, 68.2% (45) of these metabolites were each associated with the risk of a specific cancer. Among the 66 metabolites, O-methylcatechol sulfate and 4-vinylphenol sulfate demonstrated the most pronounced positive and negative associations with cancer risk, respectively. Genetically proxied plasma levels of these two metabolites were significantly associated with the risk of lung cancer and renal cell cancer, with an odds ratio and 95% confidence interval of 2.81 (2.33–3.37) and 0.49 (0.40–0.61), respectively. None of these 94 associations was biased by weak instruments, horizontal pleiotropy, or reverse causation. Further, 64 of these 94 were eligible for replication analyses, and 54 (84.4%) showed P < 0.05 with association patterns consistent with those shown in primary analyses.

*Correspondence: Weimin Li weimi003@scu.edu.cn Yaohua Yang vta8we@virginia.edu Full list of author information is available at the end of the article



© The Author(s) 2024. **Open Access** This article is licensed under a Creative Commons Attribution 4.0 International License, which permits use, sharing, adaptation, distribution and reproduction in any medium or format, as long as you give appropriate credit to the original author(s) and the source, provide a link to the Creative Commons licence, and indicate if changes were made. The images or other third party material in this article are included in the article's Creative Commons licence, unless indicated otherwise in a credit line to the material. If material is not included in the article's Creative Commons licence and your intended use is not permitted by statutory regulation or exceeds the permitted use, you will need to obtain permission directly from the copyright holder. To view a copy of this licence, visit http://creativecommons.org/licenses/by/4.0/. The Creative Commons Public Domain Dedication waiver (http://creativecommons.org/publicdomain/zero/1.0/) applies to the data made available in this article, unless otherwise stated in a credit line to the data.

Conclusions Our study unveils plausible causal relationships between 66 plasma metabolites and cancer risk, expanding our understanding of the role of circulating metabolites in cancer genetics and etiology. These findings hold promise for enhancing cancer risk assessment and prevention strategies, meriting further exploration.

Keywords Plasma metabolites, Cancer risk, Mendelian randomization

Background

Cancer is the second leading cause of human mortality, imposing substantial medical and socioeconomic burdens [1]. Consequently, the prioritization of cancer prevention and screening strategies is of critical importance. Epidemiological studies in recent decades have identified multiple genetic, lifestyle, and environmental factors associated with cancer risk [2, 3]. In particular, genomewide association studies (GWAS) have identified more than 1000 genetic susceptibility variants for various types of cancer [4–7]. However, the etiology of cancer could not be fully explained by these factors. In addition, the intricate interplay among these factors further complicates the inference of potential causality underlying their associations with cancer risk.

Circulating metabolites are small molecules originating from cells, tissues, and biological fluids, including a variety of compounds such as amino acids, carbohydrates, lipids, and xenobiotics. These molecules have been frequently employed to investigate physiological and pathophysiological processes [8, 9]. Recent observational studies have illuminated metabolic dysregulation as a hallmark of cancer, with multiple circulating metabolites linked to cancer development [10]. For example, elevated plasma pseudouridine was reported to be associated with an increased risk of ovarian cancer [11]. In addition, aberrant L-tryptophan metabolism was shown to drive the progression of breast, renal cell, and bladder cancers [12]. Therefore, investigating the metabolites associated with cancer development not only aids in early cancer screening and prevention, but also enhances our insights into the biological mechanisms underlying cancer treatment. However, these studies mainly focused on a small subset of metabolites and were limited by biases commonly encountered in conventional epidemiological studies, such as small sample sizes, potential confounders, and reverse causation.

Various factors influence metabolite levels, including genetics [13]. The advent of untargeted and targeted metabolomics technologies has facilitated the exploration of the genetic architecture of thousands of metabolites [13–17]. Typically, these investigations measure metabolite abundance in the blood, effectively reflecting the aggregative metabolite concentrations across tissues [18]. Notably, a recent study performed genotyping and untargeted plasma metabolomic profiling among 19,994 subjects of European ancestry and identified 2599 significant associations between genetic variants and metabolites [17]. Intriguingly, a considerable proportion of these variants were found to colocalize with GWAS-identified risk variants for various diseases, including cancer [17]. The shared genetic determinant over both plasma metabolites and cancer forms a strong basis for the exploration of the relationship between them using genetic variants as instrumental variables through Mendelian randomization (MR) studies. Given the random allocation of alleles during gamete formation, findings from MR analyses hold the potential to infer causal connections between exposures and outcomes [19, 20].

Several MR studies have revealed circulating metabolites with genetically predicted levels that might be causally associated with cancer risk. For example, docosapentaenoic acid [21] and high-density lipoprotein [22] were found to be associated with increased lung cancer and breast cancer risk, respectively. In additional, 1-linoleoylglycerophosphoethanolamine was associated with a reduced risk of colorectal cancer [23]. Although these results showcase the potential of circulating metabolites as causal biomarkers for cancer, it is important to note that most of these studies only investigated a limited number of metabolites, primarily owing to the relatively slow adoption of untargeted metabolomics platforms. On the other hand, a majority of these studies did not take full advantage of the most up-to-date GWAS data for both metabolites and cancers. This potential oversight could have resulted in the utilization of weak genetic instruments and less precise effect size estimates for both variant-metabolite and variant-cancer associations.

To address these limitations, we meticulously assembled the most comprehensive GWAS data available to date for untargeted metabolomics and seven cancers among individuals of European descent. Leveraging these datasets, we conducted a two-sample MR study to unravel the potential causal relationship between 913 plasma metabolites and the risk of breast, lung, colorectal, prostate, ovarian, renal cell, and testicular germ cell cancers. For the significant causal associations we identified, a series of complementary analyses were conducted to reinforce their reliability and robustness.

Methods

Study design

The overall study workflow is illustrated in Fig. 1. In this study, we treated the plasma level of each metabolite as the exposure, and the risk of each cancer as the outcome. Single nucleotide polymorphisms (SNPs) significantly associated with exposure were utilized as instrumental variables (IVs). For a robust MR study, each IV should be significantly associated with the exposure, independent of all the other IVs and potential confounding factors, and impacting the outcome only by influencing the exposure [19, 20]. These principles were carefully adhered to throughout the entire study. We conducted a comprehensive set of downstream analyses to account for potential biases that might undermine the reliability of our findings. Specifically, we estimated *F*-statistics and

conducted the Steiger test to ensure the validity of IVs, employed Egger regression and Mendelian Randomization Pleiotropy RESidual Sum and Outlier (MR-PRESSO) to detect and correct for horizontal pleiotropy and outliers, conducted leave-one-out (LOO) analyses to evaluate the presences of predominant IVs, and performed reverse MR analyses to determine the possibility of reverse causation [24-26]. We additionally implemented genetic correlation analyses to assess if shared genetic factors between metabolites and cancer risk might have confounded our MR estimates. Colocalization analyses were also conducted to examine the presence of shared causal variants between metabolites and cancer risk in genomic loci where IVs reside. For metabolites showing significant associations with the risk of each specific cancer type, we further performed multivariable MR (MVMR)



Fig. 1 Overall study design and workflow. GWAS, genome-wide association study; BC, breast cancer; CRC, colorectal cancer; RCC, renal cell cancer; LC, lung cancer; OC, ovarian cancer; PC, prostate cancer; TGCC, testicular germ cell cancer; SNPs, single nucleotide polymorphisms; kb, kilobase; MR-PRESSO, Mendelian Randomization Pleiotropy RESidual Sum and Outlier; LOO, leave-one-out

analyses [27] to nominate metabolites that might directly influence cancer risk independent of the effects of all the other metabolites. Finally, an independent plasma metabolites GWAS dataset was employed to evaluate the findings from our primary analyses.

Exposure data

GWAS data of plasma metabolites were sourced from an interactive web server accessible at https://omicscience. org/apps/mgwas/mgwas.table.php. This dataset comprised a total of 913 metabolites quantified for 14,296 individuals of European descent from the EPIC-Norfolk study and the INTERVAL study [17]. In brief, untargeted plasma metabolomic profiling was conducted using the Metabolon HD4 platform. Genotyping was performed using the Affymetrix Axiom Array, and data were imputed with the 1000 Genomes Phase 3-UK10K data as the reference panel. GWAS analyses were conducted within each cohort via linear regression analyses adjusting for age and sex, and the results were combined through inverse variance-weighted fixed-effect meta-analyses [17].

To externally validate the findings based on data from the EPIC-Norfolk and INTERVAL study, we utilized data from an independent GWAS of 1091 blood metabolites among 8192 individuals of European ancestry from the Canadian Longitudinal Study on Aging (CLSA) [14]. Summary-level GWAS statistics for these 1091 metabolites were retrieved from the GWAS catalog, under the accession numbers GCST90199621-GCST90201020.

Outcome data

Summary statistics data for GWAS on seven distinct cancers among European ancestry subjects were collected from large-scale GWAS consortia. Detailed information of these data and consortia is presented in Additional file 1: Table S1. Briefly, breast cancer data were obtained from the Breast Cancer Association Consortium (BCAC), including 142,798 cases and 123,283 controls [28]. Data on colorectal cancer were sourced from the GWAS catalog (GCST90255675), including 78,473 cases and 107,143 controls [29] from the Colorectal Cancer Transdisciplinary Study (CORECT), the Genetics and Epidemiology of Colorectal Cancer Consortium (GECCO), the Colon Cancer Family Registry (CCFR), and the UK Biobank. Data on lung cancer were obtained from the GWAS catalog (GCST004748), including 29,266 cases and 56,450 controls from the Lung Cancer Cohort Consortium (LC3) and the Transdisciplinary Research of Cancer in Lung of the International Lung Cancer Consortium (TRICL-ILCCO) [6]. Data on renal cell cancer were acquired from the database of Genotypes and Phenotypes (dbGaP; phs001736.v2.p1), including 10,784 cases and 20,406 controls from the International Agency for Research on Cancer (IARC), the National Cancer Institute (NCI), the University of Texas MD Anderson Cancer, and the Institute of Cancer Research, UK [30]. Data of prostate cancer was accessed from the Prostate Cancer Association Group to Investigate Cancer Associated Alterations in the Genome (PRACTICAL) consortium, including 79,194 cases and 61,112 controls [4]. Data on ovarian cancer were obtained from the GWAS catalog under GCST004415, including 22,406 cases and 40,941 controls from the Ovarian Cancer Association Consortium (OCAC) [31]. Data on testicular germ cell cancer were retrieved from dbGaP (phs001349.v2.p1), including 10,156 cases and 17,979 controls from the Testicular Cancer Consortium (TCC) [7].

Selection of genetic IVs

For each metabolite, non-palindromic SNPs with a minor allele frequency (MAF) of>0.05 in the 1000 Genome Project (phase 3 version 5 focusing on European descendants) and shown in cancer GWAS data were used for IV selection. Linkage disequilibrium (LD) clumping was performed with a window size of 500 kilobase (kb) to select SNPs that were independently (pairwise LD $r^2 < 0.1$) associated with plasma metabolites at $P < 1 \times 10^{-6}$, as previously described [23, 32, 33]. For each metabolite, the variance in its plasma level explained by each IV (R^2) and the strength of each IV (F-statistics) were calculated using formulas $R^2 = (2\beta^2 \times \text{EAF} \times (1 - \text{EAF}))/(2\beta^2 \times \text{EAF} \times (1 - \text{EAF})) +$ $2N \times \text{EAF} \times (1 - \text{EAF}) \times \text{SE}^2)$ and $F = (R^2 \times (N-2))/(1-R^2)$, respectively. In these formulas, EAF denotes the effect allele frequency. β and SE represent the effect size and standard error of the SNPmetabolite association, respectively. N is the sample size of the metabolite GWAS [34]. After excluding weak IVs by *F*-statistic < 10 and the Steiger test [32], metabolites with at least three IVs were eligible for MR analyses. We further applied a more stringent threshold, i.e., a window size of 1000 kb, $P < 5 \times 10^{-8}$, and LD $r^2 < 0.001$, to select IVs for MR analyses to evaluate the robustness of our findings.

MR analysis

We employed the inverse variance-weighted (IVW) method as the primary strategy for MR analyses. IVW estimates, known as the assumption of no horizontal pleiotropy across all SNPs, are derived from a comprehensive analysis of Wald ratios for all genetic variants [35]. To account for type I error, Bonferroni correction was applied to analysis results for each cancer type to identify significant associations at Bonferroni-corrected P < 0.05. To ensure the robustness of findings, complementary analyses were performed using three additional MR approaches. Specifically, the weighted median [36] method, which assumes that up to half of IVs are invalid,

was utilized to address the potential deviations from the strong assumption of IVW that all IVs are valid. Egger regression [26] was applied to identify and adjust for pleiotropic effects, wherein genetic variants influence both the exposure and the outcome. MR-PRESSO [25] was utilized to detect and correct for the impacts of outliers on MR estimates.

Complementary, sensitivity, and reverse MR analyses

To assess the robustness of significant association identified by IVW, we conducted a series of complementary and sensitivity analyses, including heterogeneity tests to assess the validity of IVs, Egger intercept test and MR-PRESSO global test to evaluate horizontal pleiotropy, and LOO analyses to examine the presence of dominant IVs [37]. To examine the possible reverse causality of the identified significant associations, we performed reverse MR analyses in which cancer was treated as the exposure and metabolites as the exposure. Given the substantially larger sample size of cancer GWAS, more stringent criteria as recommended by previous studies [32] were applied to select SNPs that were independently (pairwise LD $r^2 < 0.001$ in 1000 kb window) associated with cancer at the genome-wide statistical significance level of $P < 5 \times 10^{-8}$ as IVs. Associations with P < 0.05 estimated by the IVW method were considered significant.

Finally, a significant metabolite-cancer association was considered confident if it met a series of stringent criteria: (1) the significance of association reached Bonferronicorrected P < 0.05 using IVW as well as P < 0.05 using at least one of the other three approaches, (2) the association pattern was consistent across all MR approaches, (3) all IVs had an *F*-statistics of > 10, (4) there was no significant heterogeneity among IVs, (5) there was no evidence of horizontal pleiotropy (Egger *P* for intercept > 0.05 and MR-PRESSO global test P > 0.05), and (6) MR estimates were not significantly affected by a single IV in LOO analyses. All statistical analyses were conducted using the R packages *TwoSampleMR* (v0.5.7) [38] and *MR-PRESSO* (v1.0) [25].

Power calculation

To evaluate the statistical power of MR estimates, we utilized a specialized online tool (https://shiny.cnsge nomics.com/mRnd/) [39]. This tool employs asymptotic theory to estimate power values for the detection of causal effects derived from IVs. We performed power calculations at a type I error rate of 0.05, taking into account parameters such as R^2 of IVs, the proportion of cases of cancer GWAS, and the odds ratio (OR) of MR analyses using the IVW method.

Multivariable MR

To determine the direct impact of each plasma metabolite on cancer risk, while accounting for the effects of other metabolites, we performed multivariable MR (MVMR) analyses using the R package *MVMR* (v 0.4) [27]. MVMR effectively manages the complexities arising from interdependencies among genetic variations linked to different exposures by including multiple exposures that interact with one another [27]. For each cancer type, MVMR were performed employing all IVs involved in significant metabolite-cancer associations identified in univariate MR analyses.

Genetic correlation and colocalization analysis

MR estimates can violate causal effects in the presence of a genetic correlation between the exposure and the outcome of interest. To address this, we conducted genetic correlation analyses using linkage disequilibrium score regression (LDSC, v2.0.1), which estimates coinheritance using chi-squared statistics based on the full summary statistics of two traits [23, 40]. On the other hand, studies have suggested that colocalization analysis could complement MR by addressing its limitations related to pleiotropy and linkage disequilibrium, providing a more nuanced understanding of the shared genetic underpinnings of exposures and outcomes [41]. For each significant association identified in primary analyses, we examined the colocalization between the metabolite and the risk of cancer at each genomic locus where each IV resides using the R package coloc (v5.2.2) [42] to investigate whether identified causal associations between metabolites and each cancer risk were driven by high LD, as reported in a previous study [43]. A posterior probability (PP4) of > 0.5 was considered as evidence for moderate colocalization.

Results

Genetic IVs

We obtained summary statistics for a total of 517,882 associations between 162,261 common genetic variants (MAF>0.05) and 913 metabolites at $P < 10^{-5}$ from the study by Surendran et al. [17]. At the LD clumping criteria of pairwise LD $r^2 < 0.1$ within a 500 kb window and the significance threshold of $P < 10^{-6}$, 911 of these 913 had at least 1 IV. After excluding weak IVs based on *F*-statistics < 10 and the Steiger test, and outliers (MR-PRESSO outliner test P < 0.05), 579 metabolites, each with at least 3 IVs (median 7; interquartile range [IQR] 4–20), were retained for MR analyses. The detailed information on the IVs selected for downstream MR analyses is shown in Additional file 1: Table S2.

Overall MR results

At Bonferroni-corrected P < 0.05 for each cancer type, we identified a total of 94 significant associations, including 17 for breast cancer, 33 for colorectal cancer, 16 for lung cancer, 7 for ovarian cancer, 13 for prostate cancer, 5 for renal cell cancer, and 3 for testicular germ cell cancer (Fig. 2 and Additional file 1: Table S3). The median variance in plasma metabolite levels explained by the IVs for these associations was 10.20% (IQR 2.68-18.07%) (Additional file 1: Tables S3). Scatter plots illustrating these associations are presented in Additional file 2: Fig. S1. Out of the 66 distinct metabolites involved in these associations, 45 (68.20%) were associated with the risk of a specific cancer (Table 1), while the remaining 21 metabolites were each associated with risk of at least two different cancers (Table 2). These 66 metabolites comprised 29 lipids, 10 xenobiotics, 8 amino acids, 3 nucleotides, 1 carbohydrate, and 15 compounds that are not yet well annotated (Fig. 2). To evaluate the robustness of these significant associations, we performed MR analyses using IVs that were selected under a more stringent threshold, i.e., $P < 5 \times 10^{-8}$, LD $r^2 < 0.001$, window size of 1000 kb. Of these 94 associations, 76 had a sufficient number of IVs (> 3), 50 of which showed P < 0.05 using the IVW method. Under the Bonferroni-corrected P < 0.05, 11 of these associations remained significant (Additional file 1: Table S4).

Metabolites exclusively associated with the risk of a specific cancer

Of the 66 metabolites, 45 (68.2%) were each causally associated with specific cancer types, including 11 for breast cancer, 16 for colorectal cancer, 4 for lung cancer, 7 for ovarian cancer, 6 for prostate cancer, and 1 for renal cell cancer (Table 1). Intriguingly, none of the 7 metabolites associated with ovarian cancer exhibited any significant associations with the other 6 cancer types. In contrast, all of the three metabolites associated with testicular germ cell cancer were spontaneously associated with other cancers. Among these 45 metabolites, the strongest contributory effects on cancer risk were observed for 3-carboxy-4-methyl-5-propyl-2-furanpropanoate (CMPF) on breast cancer risk (OR 1.24; 95% confidence interval [CI] 1-(1-enyl-palmitoyl)-2-arachidonoyl-GPE 1.12 - 1.39), (P-16:0/20:4)* on colorectal cancer risk (OR 1.21; 95% CI 1.14–1.30), 1-arachidonoyl-GPE (20:4n6)* on lung cancer risk (OR 1.12; 95% CI 1.07-1.17), 4-acetamidobutanoate on ovarian cancer risk (OR 1.11; 95% CI 1.06-1.16), N6-succinyladenosine on prostate cancer risk (OR 1.21; 95% CI 1.12-1.30), and 4-guanidinobutanoate on testicular germ cell cancer risk (OR 1.14; 95% CI 1.08–1.21) (Table 1, Fig. 2). In contrast, 3,7-dimethylurate, hypotaurine, isovalerylcarnitine, 1-palmityl-GPC (O-16:0), and 1-stearoyl-GPI (18:0) showed the strongest protective effect on the risk of breast (OR 0.82; 95% CI 0.75–0.89), colorectal (OR 0.83; 95% CI 0.78–0.90), lung (OR 0.83; 95% CI 0.79–0.88), ovarian (OR 0.70; 95% CI 0.58–0.84), and prostate cancer (OR 0.78; 95% CI 0.70–0.87), respectively (Table 1, Fig. 2).

Metabolites associated with the risk of multiple cancers

Of the remaining 21 metabolites, 1 was associated with the risk of 5 cancers, 4 were each associated with the risk of 3 cancers, and 16 were each associated with the risk of 2 cancers (Table 2, Fig. 2). An unannotated metabolite X-21,410 showed significant association with the risk of most cancer types, including the increased risk of lung (OR 1.13; 95% CI 1.10-1.16) and colorectal (OR 1.29; 95% CI 1.22-1.35) cancers and decreased risk of breast (OR 0.93; 95% CI 0.91-0.94), prostate (OR 0.91; 95% CI 0.89-0.93), and renal cell (OR 0.85; 95% CI 0.81-0.90) cancers. Of the 16 metabolites each showing associations with 2 cancer types, 11 were each associated with colorectal and lung cancer risk, and notably, all of them exhibited contributory effects on both cancer types. Among them, O-methylcatechol sulfate showed the strongest association, the genetically predicted plasma levels of which were associated with a 1.41-fold (95% CI 1.31-1.52) and 2.81-fold (95% CI 2.33-3.37) increased risk of colorectal cancer and lung cancer, respectively (Table 2, Fig. 2). Conversely, 4 metabolites were each associated with both prostate and renal cell cancer, and all of them were protective against the risk of both cancers. Among them, 4-vinylphenol sulfate displayed the most pronounced protective effects, with ORs of 0.68 (95% CI 0.61-0.76) for prostate cancer and 0.49 (95% CI 0.40-0.61) for renal cell cancer (Table 2, Fig. 2).

Complementary, sensitivity, and reverse MR analyses

As shown in Additional file 1: Table S3, all the 94 significant associations identified using the IVW method consistently demonstrated the same association patterns in the results from all 3 additional MR approaches. Notably, more than 90% (85) of these associations showed P < 0.05 in at least 2 of the 3 additional MR analyses. Notably, none of these 94 associations was influenced by horizontal pleiotropy, as evidenced by Egger regression (all $P_{intercept} > 0.05$) and MR-PRESSO global test (all P > 0.05) (Additional file 1: Table S3). In addition, no significant heterogeneity was detected among IVs for any of these 94 associations (P for heterogeneity > 0.05). Further, LOO analyses confirmed that none of these 94 associations was dominated by a single IV. The statistical power



Fig. 2 Dot plot displaying IVW-based MR estimates of significant associations with Bonferroni-corrected P < 0.05 within each cancer type. OR, odds ratio. The x-axis represents seven different cancer types, while the y-axis corresponds to the identified metabolites. Circle size indicates the OR, with red indicating ORs greater than 1 and blue indicating ORs smaller than 1. A more intense red or blue color signifies associations with smaller P values

Table 1 Men	Idelian randomization (MR)	results for metabolites that were significantly associate	ed with the risk of a particu	ular cancer type using the i	inverse-variance weighted
(IVW) approad	ch				
Metabolite ID	Annotation	Number of IVs IVW	Weighted median	Egger regression	MR-PRESSO

Metabolite ID	Annotation	Number of IVs	WN			Weighted medi	an	Egger regressio	r	MR-PRESSO	
			OR (95% CI)	Р	$P_{ m bonfer}$	OR (95% CI)	Ρ	OR (95% CI)	Ρ	OR (95% CI)	Ρ
Breast cancer											
M52468	1-5tearoy -2-linoleoy -GPI (18:0/18:2)	30	0.93 (0.91–0.95)	1.67×10 ⁻⁸	7.96×10 ⁻⁶	0.94 (0.90–0.98)	7.46×10^{-3}	0.92 (0.87–0.98)	0.01	0.93 (0.91–0.95)	1.33×10 ⁻⁶
M53013	Glycosyl-N-palmitoyl-sphin- gosine	21	0.93 (0.90–0.96)	2.61×10 ⁻⁶	1.24×10 ⁻³	0.96 (0.91–1.00)	0.06	0.97 (0.91–1.04)	0.45	0.93 (0.90–0.96)	1.38×10 ⁻⁴
M46799	1-Oleoyl-3-linoleoyl-glycerol (18:1/18:2)	20	0.92 (0.89–0.95)	1.42×10 ⁻⁶	6.78×10 ⁻⁴	0.90 (0.83–0.97)	0.02	0.94 (0.82–1.07)	0.33	0.92 (0.89–0.95)	3.26×10 ⁻⁵
M36594	1-Linoleoyl-GPI (18:2)*	6	0.91 (0.87–0.95)	1.63×10^{-5}	7.75×10^{-3}	0.89 (0.83–0.95)	0.01	0.92 (0.78-1.09)	0.38	0.91 (0.87–0.95)	2.58×10^{-3}
M42574	Glycohyocholate	8	0.91 (0.87–0.95)	5.72×10^{-5}	0.03	0.94 (0.87–1.02)	0.21	0.96 (0.83-1.11)	0.60	0.91 (0.88–0.94)	1.19×10^{-3}
M20699	Erythritol	9	0.89 (0.85–0.94)	1.32×10^{-5}	6.31×10^{-3}	0.91 (0.84–0.97)	0.04	0.93 (0.85-1.02)	0.18	0.89 (0.86–0.94)	4.60×10^{-3}
M31787	3-Carboxy-4-methyl-5-propyl- 2-furanpropanoate (CMPF)	2	1.24 (1.12–1.39)	8.84×10 ⁻⁵	0.04	1.33 (1.10–1.60)	0.04	1.15 (0.69–1.94)	0.63	1.24 (1.12–1.39)	0.02
M33009	Homostachydrine*	4	0.84 (0.78–0.90)	4.37×10^{-7}	2.09×10^{-4}	0.85 (0.78–0.94)	0.04	0.95 (0.71-1.28)	0.79	0.84 (0.80-0.87)	3.21×10^{-3}
M34399	3,7-Dimethylurate	4	0.82 (0.75–0.89)	7.89×10^{-7}	3.77×10^{-4}	0.79 (0.70–0.89)	0.03	0.99 (0.62–1.57)	0.96	0.82 (0.76–0.88)	0.01
M52865	X-24544	26	0.95 (0.93-0.97)	5.62×10^{-5}	0.03	0.95 (0.91–1.00)	0.05	1.00 (0.94-1.06)	0.97	0.95 (0.93-0.97)	1.14×10^{-4}
M46398	X-21353	7	0.90 (0.86–0.95)	9.53×10^{-5}	0.04	0.91 (0.86–0.97)	0.03	0.91 (0.80-1.03)	0.20	0.90 (0.87–0.94)	2.56×10^{-3}
Colorectal can	cer										
M33387	2-Arachidonoyl-GPC (20:4)*	21	1.11 (1.08–1.13)	4.13×10^{-15}	2.07×10^{-12}	1.11 (1.08–1.14)	1.84×10^{-6}	1.11 (1.05–1.16)	5.55×10^{-4}	1.11 (1.08–1.13)	1.56×10^{-7}
M52713	1-(1-Enyl-palmitoyl)-2-palmi- toleoyl-GPC (P-16:0/16:1)*	21	1.06 (1.03–1.09)	7.60×10 ⁻⁵	0.04	1.08 (1.03–1.13)	2.99×10^{-3}	1.05 (0.99–1.12)	0.12	1.06 (1.03–1.09)	3.47×10 ⁻⁴
M01110	Arachidonate (20:4n6)	20	1.13 (1.10–1.16)	5.62×10^{-17}	2.82×10^{-14}	1.14 (1.10–1.19)	1.18×10^{-6}	1.15 (1.09–1.21)	4.41×10^{-5}	1.13 (1.10–1.16)	1.03×10^{-8}
M52613	1-(1-Enyl-stearoyl)-2-arachi- donoyl-GPC (P-18:0/20:4)	20	1.13 (1.10–1.16)	1.61×10^{-18}	8.04×10 ⁻¹⁶	1.12 (1.08–1.16)	2.98 × 10 ⁻⁶	1.13 (1.08–1.19)	1.14×10 ⁻⁴	1.13 (1.10–1.16)	6.35×10^{-9}
M00590	Hypotaurine	6	0.83 (0.78-0.90)	5.84×10^{-7}	2.93×10^{-4}	0.86 (0.75–0.99)	0.07	0.73 (0.58-0.93)	0.03	0.83 (0.79–0.88)	8.75×10^{-5}
M52673	1-(1-Enyl-palmitoyl)-2-arachi- donoyl-GPE (P-16:0/20:4)*	6	1.21 (1.14–1.30)	3.92×10 ⁻⁹	1.97×10 ⁻⁶	1.25 (1.16–1.35)	3.87 × 10 ⁻⁴	1.32 (1.16–1.49)	3.71×10^{-3}	1.21 (1.14–1.30)	3.67×10 ⁻⁴
M52715	1-(1-Enyl-palmitoyl)-2-myris- toyl-GPC (P-16:0/14:0)*	6	1.13 (1.07–1.19)	2.35×10 ⁻⁵	0.01	1.12 (1.04–1.22)	0.02	1.04 (0.90–1.21)	0.59	1.13 (1.07–1.19)	2.88×10 ⁻³
M52738	1-5tearoy -2-meadoy -GPC (18:0/20:3n9)*	6	1.16 (1.10–1.23)	1.86×10 ⁻⁷	9.31×10 ⁻⁵	1.19 (1.13–1.25)	1.88 × 10 ⁻⁴	1.21 (1.10–1.34)	7.19×10 ⁻³	1.16 (1.10–1.23)	8.09×10 ⁻⁴
M38165	Palmitoyl ethanolamide	8	1.09 (1.05–1.14)	6.08×10^{-5}	0.03	1.10 (1.03–1.16)	0.02	1.07 (0.95–1.22)	0.32	1.09 (1.06–1.13)	1.58×10^{-3}
M52474	1-(1-Enyl-palmitoyl)-GPC (P-16:0)*	7	1.19 (1.10–1.29)	2.22×10 ⁻⁵	0.01	1.22 (1.07–1.40)	0.02	0.97 (0.71–1.32)	0.85	1.19 (1.10–1.29)	4.20×10 ⁻³
M46701	X-19141	70	0.95 (0.94–0.97)	3.08×10^{-12}	1.54×10^{-9}	0.95 (0.93–0.97)	1.14×10^{-6}	0.96 (0.94–0.98)	5.96×10^{-4}	0.95 (0.94–0.96)	5.08×10^{-12}
M46601	X-11470	55	0.92 (0.90-0.95)	2.30×10^{-10}	1.15×10^{-7}	0.92 (0.86–0.98)	8.76×10^{-3}	0.93 (0.87–1.00)	0.06	0.92 (0.90–0.95)	1.70×10^{-8}
M47651	X-11491	45	0.96 (0.95–0.98)	3.72×10 ⁻⁷	1.86×10^{-4}	0.96 (0.94–0.98)	4.13×10^{-5}	0.95 (0.92-0.97)	5.86×10^{-5}	0.96 (0.95–0.98)	5.19×10 ⁻⁶

Metabolite ID	Annotation	Number of IVs	IVW			Weighted medi	an	Egger regressio	u	MR-PRESSO	
			OR (95% CI)	Р	$P_{ m bonfer}$	OR (95% CI)	Р	OR (95% CI)	Ρ	OR (95% CI)	Ρ
M47653	X-11593	34	1.04 (1.02–1.06)	1.50×10^{-5}	7.49×10 ⁻³	1.03 (1.01–1.06)	4.69×10^{-3}	1.01 (0.98–1.04)	0.48	1.04 (1.02–1.05)	1.81×10^{-6}
M46515	X-21470	24	1.06 (1.03-1.09)	7.74×10^{-5}	0.04	1.04 (1.00–1.08)	0.05	1.03 (0.98-1.09)	0.26	1.06 (1.03-1.09)	6.33×10^{-4}
M46459	X-01911	9	1.13 (1.06–1.20)	9.52×10^{-5}	0.05	1.12 (1.03-1.22)	0.04	1.15 (0.97–1.37)	0.19	1.13 (1.07–1.20)	8.61×10^{-3}
Lung cancer											
M34407	Isovalerylcarnitine	21	0.83 (0.79–0.88)	9.37×10^{-12}	4.25×10^{-9}	0.80 (0.73-0.88)	1.09×10^{-4}	0.80 (0.69-0.94)	0.01	0.83 (0.80-0.87)	5.25×10^{-8}
M35186	1-Arachidonoyl-GPE (20:4n6)*	24	1.12 (1.07–1.17)	1.32×10^{-6}	5.97×10^{-4}	1.10 (1.03-1.17)	8.96×10^{-3}	1.12 (1.02-1.22)	0.03	1.12 (1.08–1.16)	3.89×10 ⁻⁶
M15716	Imidazole lactate	44	0.95 (0.92-0.97)	7.22×10^{-5}	0.03	0.96 (0.92–0.99)	0.02	0.95 (0.90-1.00)	0.05	0.95 (0.93-0.97)	4.70×10^{-5}
M46410	X-21365	25	0.90 (0.85–0.94)	3.09×10^{-5}	0.01	0.88 (0.81–0.96)	7.63×10^{-3}	0.90 (0.80-1.00)	0.07	0.90 (0.85–0.94)	2.45×10^{-4}
Ovarian cancer											
M34401	5-Acetylamino-6-formyla- mino-3-methyluracil	41	1.07 (1.04–1.09)	8.96×10 ⁻⁷	4.36×10 ⁻⁴	1.06 (1.03–1.10)	1.77×10^{-3}	1.05 (1.01–1.10)	0.02	1.07 (1.04–1.09)	6.19×10 ⁻⁶
M34424	5-Acetylamino-6-amino- 3-methyluracil	28	1.10 (1.05–1.15)	1.71×10 ⁻⁵	8.33×10 ⁻³	1.11 (1.05–1.18)	9.77 × 10 ⁻⁴	1.12 (1.04–1.21)	6.68×10^{-3}	1.10 (1.05–1.15)	1.53×10 ⁻⁴
M01558	4-Acetamidobutanoate	25	1.11 (1.06–1.16)	2.20×10^{-5}	0.01	1.12 (1.06–1.19)	9.65×10^{-4}	1.14 (1.05–1.24)	5.08×10^{-3}	1.11 (1.06–1.16)	1.48×10^{-4}
M37496	N-acetylputrescine	25	1.10 (1.06–1.15)	1.44×10^{-5}	7.03×10^{-3}	1.11 (1.05–1.18)	1.93×10^{-3}	1.13 (1.05-1.22)	5.19×10^{-3}	1.10 (1.06–1.15)	2.87×10^{-5}
M34389	1-Methylxanthine	22	0.91 (0.87–0.95)	3.50×10^{-5}	0.02	0.89 (0.83–0.95)	1.70×10^{-3}	0.86 (0.79-0.94)	2.18×10^{-3}	0.91 (0.87-0.95)	1.90×10^{-4}
M44681	Palmitoylcarnitine	22	0.87 (0.82-0.93)	4.14×10^{-5}	0.02	0.86 (0.77–0.97)	0.02	0.82 (0.69-0.97)	0.03	0.87 (0.83-0.92)	3.25×10^{-5}
M20458	1-Palmityl-GPC (O-16:0)	5	0.70 (0.58–0.84)	9.85×10^{-5}	0.05	0.67 (0.51–0.90)	0.05	0.86 (0.43–1.74)	0.71	0.70 (0.65–0.74)	4.01×10^{-4}
Prostate cance	ir										
M18368	Cys-Gly, oxidized	36	0.94 (0.91–0.96)	5.51×10^{-6}	2.71×10^{-3}	0.92 (0.87–0.97)	3.29×10^{-3}	0.91 (0.85–0.98)	0.01	0.94 (0.91–0.96)	3.66×10^{-5}
M35136	5-Methyluridine (ribothymi- dine)	31	1.05 (1.03–1.08)	5.65×10^{-5}	0.03	1.06 (1.02–1.10)	4.71×10^{-3}	1.06 (1.01–1.12)	0.02	1.05 (1.03–1.07)	4.63×10 ⁻⁶
M33959	N-acetyltryptophan	24	0.92 (0.89–0.95)	4.15×10^{-7}	2.04×10^{-4}	0.92 (0.87–0.98)	0.01	0.94 (0.87-1.01)	0.11	0.92 (0.89–0.95)	2.30×10^{-5}
M19324	1-Stearoyl-GPI (18:0)	5	0.78 (0.70-0.87)	1.49×10^{-5}	7.33×10^{-3}	0.83 (0.68–1.01)	0.13	0.54 (0.13-2.29)	0.47	0.78 (0.70-0.87)	0.01
M48130	N6-succinyladenosine	4	1.21 (1.12–1.30)	9.48×10^{-7}	4.66×10^{-4}	1.19 (1.09–1.31)	0.03	1.22 (0.97–1.54)	0.23	1.21 (1.12–1.30)	0.02
M46471	X-12824	31	0.94 (0.91–0.97)	3.60×10^{-5}	0.02	0.94 (0.90–0.99)	0.02	0.98 (0.92–1.04)	0.43	0.94 (0.91–0.97)	2.66×10^{-4}
Renal cell cance	er										
M15681	4-Guanidinobutanoate	22	1.14 (1.08–1.21)	8.67×10^{-6}	4.12×10 ⁻³	1.12 (1.04–1.22)	8.39×10^{-3}	1.14 (1.02–1.27)	0.03	1.14 (1.08–1.20)	8.59×10 ⁻⁵
<i>IVs</i> instrumental v Mendelian Rando	variables, OR odds ratio per standar omization Pleiotropy RESidual Sum	rd deviation (SD) in g and Outlier	genetically predicte	d plasma meta	bolite level, C/ o	confidence interval,	P _{bonfer} Bonferr	oni-corrected <i>P</i> va	lue within each	i cancer type, <i>MR-H</i>	PRESSO

Table 1 (continued)

Metabolite ID	Annotation	Cancer	Number of	M			Weighted mec	lian	Egger regressic	n	MR-PRESSO	
			671	OR (95% CI)	Ρ	$P_{ m bonfer}$	OR (95% CI)	Ρ	OR (95% CI)	Ρ	OR (95% CI)	Ρ
M32827	Andro steroid monosulfate	Colorectal	38	1.10 (1.07– 1.13)	2.18×10 ⁻¹¹	1.09×10^{-8}	1.05 (1.01– 1.10)	0.03	1.06 (1.00– 1.13)	0.07	1.10 (1.07– 1.13)	7.28×10 ⁻⁸
	(1)*	Lung	22	1.20 (1.13– 1.27)	3.26×10^{-10}	1.48×10^{-7}	1.25 (1.09– 1.43)	4.22×10^{-3}	1.22 (1.02– 1.45)	0.04	1.20 (1.13– 1.27)	3.11×10 ⁻⁶
		Testicular germ cell	53	0.86 (0.80– 0.91)	1.04 × 10 ⁻⁶	5.00×10^{-4}	0.86 (0.78– 0.96)	9.26×10^{-3}	0.88 (0.77– 1.00)	0.05	0.86 (0.80– 0.91)	1.03×10 ⁻⁵
M52689	1-(1-Enyl- palmitoyl)-	Colorectal	32	1.11 (1.09– 1.14)	2.92×10^{-21}	1.46 × 10 ⁻¹⁸	1.12 (1.08– 1.15)	1.21×10^{-7}	1.12 (1.07– 1.16)	1.14×10 ⁻⁵	1.11 (1.09– 1.14)	9.51×10^{-11}
	2-arachi- donoyl-GPC (P-16:0/20:4)*	Lung	29	1.12 (1.08– 1.16)	7.29×10 ⁻⁹	3.31×10 ⁻⁶	1.10 (1.03– 1.16)	4.10×10 ⁻³	1.10 (1.03– 1.17)	9.51×10^{-3}	1.12 (1.08– 1.15)	1.29×10 ⁻⁷
M33228	1-Arachi- donoyl-GPC	Colorectal	26	1.09 (1.07– 1.11)	1.95 × 10 ⁻¹⁵	9.78×10 ⁻¹³	1.10 (1.07– 1.13)	7.19×10 ⁻⁷	1.09 (1.05– 1.13)	4.32×10 ⁻⁵	1.09 (1.07– 1.11)	4.17×10 ⁻⁹
	(20.4n6)*	Lung	27	1.10 (1.06– 1.13)	3.49×10 ⁻⁸	1.58×10^{-5}	1.08 (1.04– 1.14)	1.86×10^{-3}	1.09 (1.04– 1.15)	3.11×10^{-3}	1.10 (1.06– 1.13)	7.38×10 ⁻⁶
M52462	1-Palmitoyl- 2-arachi-	Colorectal	27	1.10 (1.08– 1.13)	3.27×10^{-18}	1.64×10^{-15}	1.10 (1.07– 1.13)	3.26×10^{-7}	1.10 (1.06– 1.14)	5.62×10^{-5}	1.10 (1.08– 1.13)	3.53×10^{-9}
	donoyl-GPC (16:0/20:4)	Lung	27	1.10 (1.07– 1.14)	5.56×10^{-8}	2.52×10^{-5}	1.09 (1.05– 1.14)	2.76×10 ⁻⁴	1.10 (1.03– 1.17)	6.29×10^{-3}	1.10 (1.07– 1.14)	1.08×10 ⁻⁵
M42450	1-Stearoyl- 2-arachi-	Colorectal	25	1.09 (1.07– 1.11)	6.62×10^{-19}	3.32×10^{-16}	1.09 (1.06– 1.12)	1.05×10^{-6}	1.08 (1.05– 1.12)	8.22×10^{-5}	1.09 (1.07– 1.11)	4.29×10 ⁻⁹
	donoyl-GPC (18:0/20:4)	Lung	25	1.09 (1.06– 1.13)	1.52×10^{-8}	6.91 × 10 ⁻⁶	1.08 (1.04– 1.13)	8.84×10^{-4}	1.07 (1.02– 1.13)	0.01	1.09 (1.06– 1.13)	6.71×10 ⁻⁶
M52719	Phosphati- dylcholine	Colorectal	21	1.14 (1.10– 1.17)	1.03×10^{-17}	5.16×10^{-15}	1.15 (1.10– 1.19)	1.24×10 ⁻⁶	1.16 (1.10– 1.22)	2.40×10^{-5}	1.14 (1.10– 1.17)	3.96×10 ⁻⁸
	(16:0/22:5n3, 18:1/20:4)*	Lung	23	1.10 (1.06– 1.16)	1.25×10^{-5}	5.66×10^{-3}	1.11 (1.04– 1.18)	3.61×10^{-3}	1.13 (1.05– 1.23)	6.44×10^{-3}	1.10 (1.06– 1.15)	2.11×10 ⁻⁴
M52698	1-Adrenoyl- GPC (22:4)*	Colorectal	6	1.20 (1.14– 1.27)	6.02×10^{-11}	3.01 × 10 ⁻⁸	1.25 (1.17– 1.34)	2.01 × 10 ⁻⁴	1.30 (1.16– 1.44)	2.14×10^{-3}	1.20 (1.14– 1.27)	1.80×10 ⁻⁴
		Lung	6	1.17 (1.08– 1.27)	6.76×10 ⁻⁵	0.03	1.17 (1.05– 1.29)	0.02	1.21 (0.99– 1.48)	0.10	1.17 (1.09– 1.27)	3.99×10 ⁻³
M46111	O-methylcat- echol sulfate	Colorectal	6	1.41 (1.31– 1.52)	6.25×10^{-19}	3.13×10^{-16}	1.49 (1.25– 1.78)	2.35×10^{-3}	1.87 (1.34– 2.62)	8.21×10^{-3}	1.41 (1.31– 1.51)	1.18×10 ⁻⁵
		Lung	4	2.81 (2.33– 3.37)	4.32×10 ⁻²⁸	1.96×10^{-25}	2.85 (2.04– 4.00)	8.86×10^{-3}	4.03 (1.35– 12.06)	0.13	2.81 (2.48– 3.17)	4.82×10 ⁻⁴
M34397	1-Arachidonyl- glycerol (20:4)	Colorectal	Q	1.18 (1.13– 1.24)	9.69×10 ⁻¹³	4.86×10^{-10}	1.20 (1.14– 1.27)	1.50×10^{-3}	1.28 (1.13– 1.45)	0.02	1.18 (1.13– 1.23)	4.19×10 ⁻⁴
		Lung	œ	1.19 (1.11– 1.27)	7.87×10^{-7}	3.57 × 10 ⁻⁴	1.14 (1.05– 1.24)	0.02	1.07 (0.93– 1.23)	0.40	1.19 (1.13– 1.25)	4.07×10 ⁻⁴

Table 2 MR results for metabolites that were significantly associated with the risk of more than one cancer type using the IVW approach

Metabolite ID	Annotation	Cancer	Number of	IVW			Weighted med	dian	Egger regressi	uo	MR-PRESSO	
			IVS	OR (95% CI)	Р	P_{bonfer}	OR (95% CI)	Ρ	OR (95% CI)	Α	OR (95% CI)	Р
M35157	N6-carbamoyl- threonyladen-	Colorectal	9	1.34 (1.21– 1.49)	1.83 × 10 ⁻⁸	9.19×10 ⁻⁶	1.24 (1.00– 1.54)	0.11	1.97 (1.19– 3.27)	0.06	1.34 (1.22– 1.48)	1.65×10 ⁻³
	osine	Lung	5	1.71 (1.47– 2.00)	1.38×10^{-11}	6.28×10^{-9}	1.61 (1.23– 2.12)	0.03	2.77 (1.36– 5.66)	0.07	1.71 (1.48– 1.99)	2.11×10^{-3}
		Breast	7	0.75 (0.68– 0.82)	3.81×10^{-9}	1.82 × 10 ⁻⁶	0.84 (0.73– 0.97)	0.05	0.65 (0.41– 1.03)	0.13	0.75 (0.68– 0.82)	1.06×10^{-3}
M46455	X-21410	Colorectal	45	1.13 (1.10– 1.16)	3.61×10^{-19}	1.81×10 ⁻¹⁶	1.15 (1.08– 1.24)	1.97×10 ⁻⁴	1.19 (1.12– 1.26)	3.15×10^{-6}	1.13 (1.10– 1.16)	1.83×10 ⁻¹¹
		Lung	25	1.29 (1.22– 1.35)	7.06×10^{-23}	3.21×10^{-20}	1.24 (1.14– 1.35)	3.79×10 ⁻⁵	1.17 (1.05– 1.30)	0.01	1.29 (1.22– 1.35)	6.64×10^{-10}
		Breast	45	0.93 (0.91– 0.94)	5.24×10^{-14}	2.50×10^{-11}	0.93 (0.90– 0.97)	2.44×10^{-3}	0.94 (0.90– 0.99)	0.01	0.93 (0.91– 0.94)	1.95×10^{-9}
		Prostate	46	0.91 (0.89– 0.93)	5.48×10^{-12}	2.70 × 10 ⁻⁹	0.89 (0.84– 0.95)	6.35×10^{-4}	0.88 (0.83– 0.94)	3.37×10 ⁻⁴	0.91 (0.89– 0.93)	1.47×10 ⁻⁸
		Renal cell	46	0.85 (0.81– 0.90)	1.19×10 ⁻⁹	5.66×10^{-7}	0.94 (0.85– 1.04)	0.24	0.89 (0.80– 0.99)	0.04	0.85 (0.81– 0.90)	1.62×10^{-7}
M36098	4-Vinylphenol sulfate	Prostate	8	0.68 (0.61– 0.76)	9.77×10^{-13}	4.81 × 10 ⁻¹⁰	0.74 (0.62– 0.90)	0.02	0.52 (0.31– 0.88)	0.05	0.68 (0.61– 0.76)	1.88×10 ⁻⁴
		Renal cell	80	0.49 (0.40– 0.61)	1.77×10^{-11}	8.43 × 10 ⁻⁹	0.41 (0.24– 0.71)	0.01	0.38 (0.04– 3.57)	0.43	0.49 (0.41– 0.59)	1.28×10 ⁻⁴
M37181	4-Allylphenol sulfate	Prostate	8	0.68 (0.63– 0.74)	9.30×10^{-19}	4.58×10^{-16}	0.64 (0.52– 0.79)	4.39×10 ⁻³	0.88 (0.55– 1.39)	0.60	0.68 (0.63– 0.74)	4.78×10 ⁻⁵
		Renal cell	6	0.54 (0.45– 0.64)	8.33×10^{-13}	3.96×10^{-10}	0.68 (0.49– 0.95)	0.05	1.35 (0.45– 4.09)	0.61	0.54 (0.45– 0.64)	9.66×10 ⁻⁵
		Colorectal	5	1.32 (1.18– 1.48)	2.11×10 ⁻⁶	1.06×10^{-3}	1.41 (1.15– 1.73)	0.03	1.13 (0.49– 2.61)	0.80	1.32 (1.21– 1.44)	3.56×10^{-3}
M46356	X-21312	Prostate	25	0.85 (0.82– 0.89)	3.84×10^{-15}	1.89×10 ⁻¹²	0.91 (0.85– 0.97)	0.01	0.79 (0.71– 0.88)	3.23×10^{-4}	0.85 (0.82– 0.89)	4.31×10^{-8}
		Renal cell	27	0.67 (0.62– 0.72)	4.00×10^{-23}	1.90×10^{-20}	0.63 (0.52– 0.77)	1.33×10 ⁻⁴	0.57 (0.42– 0.75)	6.79×10 ⁻⁴	0.67 (0.62– 0.72)	1.18×10 ⁻¹¹
		Breast	24	0.86 (0.83– 0.88)	5.71×10^{-25}	2.72×10^{-22}	0.90 (0.83– 0.98)	0.03	0.82 (0.73– 0.92)	2.17×10 ⁻³	0.86 (0.83– 0.88)	4.20×10 ⁻¹⁰
M38168	16a-Hydroxy DHEA 3-sulfate	Colorectal	30	1.06 (1.03– 1.08)	6.26×10^{-5}	0.03	1.05 (1.01– 1.09)	0.03	1.04 (0.98– 1.10)	0.19	1.06 (1.03– 1.08)	1.01×10 ⁻⁴
		Testicular germ cell	1 42	0.88 (0.82– 0.94)	7.00×10^{-5}	0.03	0.89 (0.82– 0.97)	0.01	0.88 (0.79– 0.98)	0.02	0.88 (0.82– 0.94)	2.78×10 ⁻⁴

Table 2 (continued)

Metabolite ID	Annotation	Cancer	Number of	M			Weighted med	lian	Egger regression	u	MR-PRESSO	
			5	OR (95% CI)	Р	$P_{ m bonfer}$	OR (95% CI)	Р	OR (95% CI)	Ρ	OR (95% CI)	Р
M27719	Galactonate	Colorectal	2	0.81 (0.76– 0.87)	3.67 × 10 ⁻⁹	1.84 × 10 ⁻⁶	0.81 (0.73– 0.90)	0.02	0.82 (0.64– 1.04)	0.19	0.81 (0.78– 0.85)	4.64×10 ⁻⁴
		Testicular germ cell	5	0.64 (0.52– 0.78)	9.60×10 ⁻⁶	4.62 × 10 ⁻³	0.66 (0.49– 0.88)	0.05	0.93 (0.46– 1.86)	0.84	0.64 (0.57– 0.71)	1.52×10^{-3}
M52449	1-Stearoyl- 2-arachi-	Breast	22	1.06 (1.03– 1.08)	8.31 × 10 ⁻⁵	0.04	1.09 (1.03– 1.15)	4.44×10^{-3}	1.05 (0.99– 1.12)	0.13	1.06 (1.03– 1.08)	5.25×10 ⁻⁴
	donoyl-GPI (18:0/20:4)	Colorectal	22	1.09 (1.05– 1.13)	2.27 × 10 ⁻⁵	0.01	1.06 (1.00– 1.12)	0.06	1.07 (0.97– 1.18)	0.20	1.09 (1.05– 1.13)	3.69×10 ⁻⁴
M36099	4-Ethylphenyl- sulfate	Colorectal	Ø	1.42 (1.30– 1.55)	5.07×10^{-14}	2.54×10^{-11}	1.50 (1.26– 1.79)	2.66×10^{-3}	1.52 (1.00– 2.31)	0.10	1.42 (1.30– 1.55)	1.34×10 ⁻⁴
		Prostate	8	0.69 (0.63– 0.75)	9.05 × 10 ⁻¹⁶	4.45×10^{-13}	0.65 (0.54– 0.77)	1.91×10^{-3}	0.76 (0.45– 1.27)	0.34	0.69 (0.63– 0.75)	8.83×10 ⁻⁵
M46294	X-21285	Breast	24	0.92 (0.89– 0.95)	1.95×10^{-7}	9.30 × 10 ⁻⁵	0.86 (0.77– 0.96)	0.01	0.95 (0.82– 1.11)	0.55	0.92 (0.89– 0.95)	2.82×10 ⁻⁵
		Colorectal	25	1.12 (1.08– 1.16)	8.72×10^{-10}	4.37×10^{-7}	1.16 (1.06– 1.26)	2.40×10^{-3}	1.20 (1.07– 1.34)	4.08×10^{-3}	1.12 (1.08– 1.16)	2.47×10 ⁻⁶
M46590	X-07765	Lung	4	1.47 (1.25– 1.73)	3.41 × 10 ⁻⁶	1.55×10^{-3}	1.36 (1.15– 1.61)	0.04	1.55 (0.84– 2.86)	0.30	1.47 (1.25– 1.73)	0.02
		Prostate	ø	0.81 (0.75– 0.87)	6.92×10^{-8}	3.41 × 10 ⁻⁵	0.85 (0.78– 0.93)	0.01	0.73 (0.56– 0.95)	0.06	0.81 (0.75– 0.87)	1.02×10 ⁻³
M46710	X-17690	Breast	9	0.86 (0.80– 0.92)	1.06×10 ⁻⁵	5.06×10^{-3}	0.89 (0.79– 0.99)	0.08	0.88 (0.58– 1.31)	0.56	0.86 (0.80– 0.92)	6.99×10 ⁻³
		Prostate	6	0.79 (0.73– 0.86)	1.69×10^{-8}	8.33 × 10 ⁻⁶	0.85 (0.78– 0.93)	6.11×10^{-3}	0.91 (0.64– 1.27)	0.59	0.79 (0.73– 0.86)	4.87×10 ⁻⁴
<i>IVs</i> instrumental v Pleiotropy RESidu	rariables, OR odds I	ratio per SD in gen. r	etically predicted pl	asma metabolite l	evel, <i>Cl</i> confider	interval, P _{bor}	_{nér} Bonferroni-co	rrected <i>P</i> value	within each cance	r type, <i>MR-PRE</i>	SSO Mendelian Ra	ndomization

Chen et al. BMC Medicine (2024) 22:90

Table 2 (continued)

of all MR estimates based on the IVW methods ranged from 0.98 to 1.00 (Additional file 1: Table S3). In reverse MR analyses, 195, 133, 15, 12, 158, 12, and 48 SNPs were selected as IVs for breast, colorectal, lung, ovarian, prostate, renal cell, and testicular germ cell cancers, respectively (Additional file 1: Table S5). Utilizing these IVs and the IVW method, none of the 94 associations showed the possibility of reverse causation (all P>0.05) (Additional file 1: Table S3).

Replication analysis using an independent plasma metabolite GWAS dataset

Of the 64 unique metabolites included in the 94 significant associations, 44 metabolites in 64 associations had available data in the CLSA study. After selecting IVs using the same criteria, MR analyses were performed using the same cancer GWAS data with IVW as the primary method, supplemented by weighted median, Egger regression, and MR-PRESSO. Remarkably, all but 2 of these 64 associations showed an association direction that is consistent with those observed analyses using EPIC-Norfold and INTERVAL data, and 54 of them (84.4%) even reached the nominal significance of P < 0.05(Additional file 1: Table S6).

MVMR analyses

To uncover metabolites that might be directly associated with cancer risk independent of other metabolites, we conducted MVMR analyses. For each cancer type, IVs for all significant associations identified in univariate MR analyses were included in MVMR analyses using the IVW model. A total of 21 metabolites were found to be independently associated with cancer risk at MVMR P < 0.05, including 7 associated with breast cancer risk, 5 associated colorectal cancer risk, 2 associated with lung cancer risk, 3 associated with ovarian cancer risk, 3 associated with prostate cancer risk, and 1 associated with testicular germ cell cancer risk (Additional file 1: Table S7).

Genetic correlation and colocalization analyses

Among the 94 identified significant associations, LDSC analyses detected nominally significant (P < 0.05) genetic correlations between seven metabolite-cancer pairs (Additional file 1: Table S8). This result suggests that for most of our identified significant associations, the causal effects were unlikely to be confused by the coheritability between metabolites and cancer risk. On the other hand, in colocalization analyses, 70 (74.5%) of these 94 metabolite-cancer pairs showed a moderate colocalization with (PP4>0.5) in at least 1 locus where their IVs reside (Additional file 1: Table S9), indicating the existence of shared causal variants between metabolites and cancer risk in these genomic regions. Among these 70

metabolite-cancer pairs, the median percentage of IVs whose loci exhibited colocalization signals was 50.0% (IQR 20.0–68.1%).

Discussion

In this comprehensive MR study empowered by the unprecedented resources of large-scale GWAS data, we discovered 94 significant associations indicating the potential causal influences of 66 unique plasma metabolites on the risk of 7 cancers. Over two-thirds of these metabolites were exclusively identified for specific cancer types. Of the 64 associations eligible for external validation analyses, nearly 85% (54) were successfully replicated. Further, MVMR analyses revealed that 21 of these 66 metabolites likely have direct effects on cancer risk. These findings provide additional insights into the complex interplay between genetics and metabolites in cancer development, fostering the development of innovative strategies for cancer prevention and treatment.

Developing effective strategies for cancer risk assessment and prevention is critically important. The emergence of metabolomics technologies has fueled interest in exploring the clinical utility of circulating metabolites as a non-invasive biomarker, given their ability to reflect both endogenous and exogenous physiological processes [18, 44]. Metabolic molecules, such as those involved in nucleotide metabolism, have shown potential as therapeutic targets in impeding tumor progression in preclinical studies [45]. Although previous studies have identified metabolites involved in cancer mechanisms, their role in risk assessment and prevention was constrained by unclear causal links [46]. Leveraging recent large metabolite GWAS data, well-powered cancer GWAS data, and the MR framework, we systematically explored the potential causal relationship between plasma metabolites and the risk of common cancers. Our findings, if future validated by future case-control studies nested in large population-based cohorts, hold the potential to significantly contribute to the development of metabolites-based panels for cancer risk stratification and the identification of new therapeutic targets, thereby substantially improving cancer management and treatment strategies.

Although the majority of the 94 significant metabolite-cancer associations were first reported by our study, several of them are in line with the findings from previous studies. The negative association between isovalerylcarnitine, a specific activator of high calcium [47], and lung cancer risk is consistent with a recent report based on both MR and nested case-control investigations. The 4-guanidinobutanoate, an intermediate product in the polyamine synthesis pathway, was found to be associated with increased renal cell cancer risk in our study. This metabolite was previously reported to be correlated with an increased estimated glomerular filtration rate (eGFR), indicating its possible role in kidney dysfunction [48]. Oxidized Cys-Gly, which showed protective effects on prostate cancer risk in the current study, was previously found to be associated with a decreased risk of gastric cardia adenocarcinoma [49]. All of these showcase the validity of our findings.

We found 22 metabolites each showing associations with more than one cancer type. Interestingly, 11 of them were spontaneously associated with the increased risk of both lung cancer and colorectal cancer. These results might be partially explained by their shared metabolicrelated risk factors, such as physical inactivity [50, 51] and a diet low in fiber [52, 53]. Future studies are needed to appraise the putative shared genetic and metabolic architecture of these two cancers. On the other hand, some metabolites showed contradictory effects on different cancers, such as N6-carbamoylthreonyladenosine, which was linked to an elevated risk of lung cancer and colorectal cancer but a reduced risk of breast cancer. This metabolite was previously correlated with elevated blood interleukin-6 in older adults, which was associated with an increased risk of cancer and mortality [54]. The possible protective effects of this metabolite on breast cancer need further investigation.

The human metabolome is profoundly influenced by a wide range of endogenous and exogenous factors, including genetic as well as dietary-, drug-, and diseaserelated influences, making etiologic studies interrogating its impacts on various cancers extremely difficult. MR largely overcomes those by relying upon the random assignment of alleles at conception, yet it can yield unbiased causal estimates when its assumptions are strictly followed [19]. In addition, we utilized data from the largest GWAS of untargeted plasma metabolome and cancers to date, which ensured unparalleled statistical power for selecting robust IVs with high-accuracy association estimates for MR analyses. Furthermore, a series of complementary analyses were performed to strengthen the reliability and robustness of the findings, including different MR approaches to account for the potential violation of different MR assumptions, LOO analyses to detect associations driven by a single IV, and reverse MR to assess the possibility of reverse causation. Finally, nearly 85% of significant associations that were eligible for external validation were successfully replicated, highlighting the robustness of our findings.

Our study has limitations. First, both metabolites and cancer GWASs focused on individuals of European ancestry due to the small sample sizes of such datasets in understudied populations. This hampered the evaluation of racial/ethnic disparities in metabolite-cancer associations. Second, using sex-specific GWAS for metabolites is ideal for sex-specific cancers; however, such data were not released by either the EPIC-Norfolk and INTERVAL study or the CLSA study [14, 17]. Then, investigating metabolites in cancer-relevant normal tissues would provide more etiological insights. However, metabolomic profiling of solid tissues remains a challenging task. It is well-recognized that for many metabolites, plasma levels represent the aggregation of tissue levels [18]. Therefore, the associations observed in plasma-based analyses should at least partially reflect the carcinogenic roles of these metabolites in tissues. Further, populationbased cohort studies of measured plasma metabolite levels and cancer risk, as well as in vitro investigations of the functions of metabolites in cell lines or animal models, are ideal to validate our findings. However, we were unable to carry out such studies due to the unavailability of related resources.

Conclusions

In this systemic MR study, we unveiled compelling evidence supporting putative causal links between 66 plasma metabolites and the risk of seven cancers, a large proportion of which were successfully replicated. Our results contribute to an advanced understanding of the crucial role of circulating metabolites in cancer genetics and biology. The utility of these metabolites in cancer risk assessment and prevention merits further investigation.

Abbreviations

BCAC	Breast Cancer Association Consortium
CCFR	Colon Cancer Family Registry
CLSA	Canadian Longitudinal Study on Aging
CORECT	Colorectal Cancer Transdisciplinary Study
GECCO	Genetics and Epidemiology of Colorectal Cancer Consortium
GWAS	Genome-wide association study
IV	Instrumental variable
IVW	Inverse variance-weighted
LC3	Lung Cancer Cohort Consortium
LD	Linkage disequilibrium
LDSC	Linkage disequilibrium score regression
LOO	Leave-one-out
MAF	Minor allele frequency
MR	Mendelian randomization
MR-PRESSO	Mendelian Randomization Pleiotropy RESidual Sum and Outlier
MVMR	Multivariable Mendelian randomization
OCAC	Ovarian Cancer Association Consortium
PP	Posterior probability
PRACTICAL	Prostate Cancer Association Group to Investigate Cancer Asso-
	ciated Alterations in the Genome
TCC	Testicular Cancer Consortium
TRICL-ILCCO	Transdisciplinary Research of Cancer in Lung of the Interna- tional Lung Cancer Consortium
	-

Supplementary Information

The online version contains supplementary material available at https://doi. org/10.1186/s12916-024-03272-8.

Additional file 1: Table S1. Information on genome-wide association studies (GWAS) of cancers among individuals of European ancestry. Table S2. Summary statistics of instrumental variables (IVs) for the 94 significant associations identified in primary Mendelian randomization (MR) analyses. Table S3. Significant associations identified in primary MR analyses using the inverse-variance weighted (IVW) method. Table S4. MR results based on instrumental variables (IVs) selected under a more stringent threshold ($P < 5 \times 10^{-8} \& LD r^2 < 0.001$) for the 94 significant associations identified in primary analyses. Table S5. Summary statistics of instrumental variables (IVs) used for reverse Mendelian randomization (MR) analyses for the 94 significant associations identified in primary analyses. Table S6. Replication analyses results using an independent metabolite GWAS dataset for the 94 significant associations identified in primary analyses. Table S7. Twenty-one associations that remained significant in Multivariable MR (MVMR) analyses. Table S8. Seven metabolitecancer pairs showing a nominally significant genetic correlation. Table S9. Colocalization analyses for the 94 significant associations identified in primary analyses.

Additional file 2: Fig. S1. Scatter plots of the identified 94 significant metabolite-cancer associations.

Acknowledgements

The acknowledgments to GWAS consortia were described in detail in Additional file 1: Table S1. Data analyses were partially conducted on the Rivanna High-Performance Computing system at the University of Virginia.

Authors' contributions

YY and WL: conceptualization, supervision, validation, investigation, methodology, writing—review, and editing. YY, WL, and YC: funding acquisition. YC: project administration, formal analysis, visualization, writing—original draft, and writing—review. ZC and YK: methodology and data curation. HC and SL: visualization, validation, writing—review, and editing. JT, YX, and DL: investigation, writing—review, and editing. GW and YQ: data curation and writing review. All authors contributed to the planning, execution, and analysis of the study and reviewed and approved the final submitted version.

Author's Twitter handles

@YaohuaYang_1989 (Yaohua Yang).

Funding

Y. Yang was supported partially by the National Cancer Institute (No. R00CA248822). W. Li was supported by the National Natural Science Foundation of China (No. 92159302). Y. Chen was supported by the National Natural Science Foundation of China (No. 82300011) and the Sichuan Science and Technology Support Project (No. 2022NSFSC1516). The sponsors had no role in the study design; in the collection, analysis, or interpretation of the data; in the writing of the report; or in the decision to submit the paper for publication.

Availability of data and materials

All used in the present study are publicly available. The original repositories of these data are described in the "Methods" section and Additional file 1: Table S1. Data that could be used to replication our findings is available in Additional file 1: Table S2 and S5.

Declarations

Ethics approval and consent to participate

All the data used in this study were acquired from publicly available genomewide association study summary statistics. No new data was collected, and no new ethical approval was needed.

Consent for publication

Not applicable.

Competing interests

The authors declare that they have no competing interests.

Author details

¹Institute of Respiratory Health, Frontiers Science Center for Disease-Related Molecular Network, West China Hospital, Sichuan University, Guoxue Alley 37, Chengdu, Sichuan, China. ²Department of Respiratory and Critical Care Medicine, West China Hospital, Sichuan University, Chengdu, Sichuan, China. ³Center for Public Health Genomics, Department of Public Health Sciences, UVA Comprehensive Cancer Center, School of Medicine, University of Virginia, 560 Ray C. Hunt Dr., Rm 4408, Charlottesville, VA, USA. ⁴Department of Public Health Sciences, UVA Comprehensive Cancer Center, School of Medicine, University of Virginia, Charlottesville, VA, USA. ⁴Department of Public Health Sciences, UVA Comprehensive Cancer Center, School of Medicine, University of Virginia, Charlottesville, VA, USA. ⁵Innovation Laboratory for Precision Diagnostics, Precision Medicine Research Center, Precision Medicine Key Laboratory of Sichuan Province, West China Hospital, Sichuan University, Chengdu, Sichuan, China.

Received: 7 September 2023 Accepted: 22 January 2024 Published online: 04 March 2024

References

- Siegel RL, Miller KD, Wagle NS, Jemal A. Cancer statistics, 2023. CA-Cancer J Clin. 2023;73(1):17–48.
- 2. Cao M, Li H, Sun D, He S, Yan X, Yang F, et al. Current cancer burden in China: epidemiology, etiology, and prevention. Cancer Biol Med. 2022;19(8):1121.
- Thrift AP, Wenker TN, El-Serag HB. Global burden of gastric cancer: epidemiological trends, risk factors, screening and prevention. Nat Rev Clin Oncol. 2023;20(5):338–49.
- Conti DV, Darst BF, Moss LC, Saunders EJ, Sheng X, Chou A, Schumacher FR, Olama AAA, Benlloch S, Dadaev T, et al. Trans-ancestry genome-wide association meta-analysis of prostate cancer identifies new susceptibility loci and informs genetic risk prediction. Nat Genet. 2021;53(1):65–75.
- Fernandez-Rozadilla C, Timofeeva M, Chen Z, Law P, Thomas M, Schmit S, Díez-Obrero V, Hsu L, Fernandez-Tajes J, Palles C. Deciphering colorectal cancer genetics through multi-omic analysis of 100,204 cases and 154,587 controls of European and east Asian ancestries. Nat Genet. 2023;55(1):89–99.
- Byun J, Han Y, Li Y, Xia J, Long E, Choi J, Xiao X, Zhu M, Zhou W, Sun R, et al. Cross-ancestry genome-wide meta-analysis of 61,047 cases and 947,237 controls identifies new susceptibility loci contributing to lung cancer. Nat Genet. 2022;54(8):1167–77.
- Pluta J, Pyle LC, Nead KT, Wilf R, Li M, Mitra N, Weathers B, D'Andrea K, Almstrup K, Anson-Cartwright L, et al. Identification of 22 susceptibility loci associated with testicular germ cell tumors. Nat Commun. 2021;12(1):4487.
- Wishart DS. Metabolomics for investigating physiological and pathophysiological processes. Physiol Rev. 2019;99(4):1819–75.
- van der Spek A, Stewart ID, Kühnel B, Pietzner M, Alshehri T, Gauß F, Hysi PG, MahmoudianDehkordi S, Heinken A, Luik AI. Circulating metabolites modulated by diet are associated with depression. Mol Psychiatry. 2023;28(9):3874–87.
- 10. Pavlova NN, Zhu J, Thompson CB. The hallmarks of cancer metabolism: still emerging. Cell Metab. 2022;34(3):355–77.
- Zeleznik OA, Eliassen AH, Kraft P, Poole EM, Rosner BA, Jeanfavre S, Deik AA, Bullock K, Hitchcock DS, Avila-Pacheco J, et al. A prospective analysis of circulating plasma metabolites associated with ovarian cancer risk. Cancer Res. 2020;80(6):1357–67.
- Platten M, Nollen EAA, Rohrig UF, Fallarino F, Opitz CA. Tryptophan metabolism as a common therapeutic target in cancer, neurodegeneration and beyond. Nat Rev Drug Discov. 2019;18(5):379–401.
- Shin SY, Fauman EB, Petersen AK, Krumsiek J, Santos R, Huang J, Arnold M, Erte I, Forgetta V, Yang TP, et al. An atlas of genetic influences on human blood metabolites. Nat Genet. 2014;46(6):543–50.
- Chen Y, Lu T, Pettersson-Kymmer U, Stewart ID, Butler-Laporte G, Nakanishi T, Cerani A, Liang KYH, Yoshiji S, Willett JDS, et al. Genomic atlas of the plasma metabolome prioritizes metabolites implicated in human diseases. Nat Genet. 2023;55(1):44–53.
- 15. Yin X, Chan LS, Bose D, Jackson AU, VandeHaar P, Locke AE, Fuchsberger C, Stringham HM, Welch R, Yu K, et al. Genome-wide association studies of

metabolites in Finnish men identify disease-relevant loci. Nat Commun. 2022;13(1):1644.

- Kettunen J, Demirkan A, Wurtz P, Draisma HH, Haller T, Rawal R, Vaarhorst A, Kangas AJ, Lyytikainen LP, Pirinen M, et al. Genome-wide study for circulating metabolites identifies 62 loci and reveals novel systemic effects of LPA. Nat Commun. 2016;7:11122.
- 17. Surendran P, Stewart ID, Au Yeung VPW, Pietzner M, Raffler J, Worheide MA, Li C, Smith RF, Wittemans LBL, Bomba L, et al. Rare and common genetic determinants of metabolic individuality and their effects on human health. Nat Med. 2022;28(11):2321–32.
- Bartel J, Krumsiek J, Schramm K, Adamski J, Gieger C, Herder C, Carstensen M, Peters A, Rathmann W, Roden M. The human blood metabolome-transcriptome interface. PLoS Genet. 2015;11(6):e1005274.
- 19. Emdin CA, Khera AV, Kathiresan S. Mendelian randomization. JAMA. 2017;318(19):1925–6.
- Boef AG, Dekkers OM, le Cessie S. Mendelian randomization studies: a review of the approaches used and the quality of reporting. Int J Epidemiol. 2015;44(2):496–511.
- Zhao H, Wu S, Luo Z, Liu H, Sun J, Jin X. The association between circulating docosahexaenoic acid and lung cancer: a Mendelian randomization study. Clin Nutr. 2022;41(11):2529–36.
- Liu J, Zhou H, Zhang Y, Huang Y, Fang W, Yang Y, Hong S, Chen G, Zhao S, Chen X, et al. Docosapentaenoic acid and lung cancer risk: a Mendelian randomization study. Cancer Med. 2019;8(4):1817–25.
- Yun Z, Guo Z, Li X, Shen Y, Nan M, Dong Q, Hou L. Genetically predicted 486 blood metabolites in relation to risk of colorectal cancer: a Mendelian randomization study. Cancer Med. 2023;12(12):13784–99.
- Burgess S, Thompson SG. Bias in causal estimates from Mendelian randomization studies with weak instruments. Stat Med. 2011;30(11):1312–23.
- Verbanck M, Chen CY, Neale B, Do R. Detection of widespread horizontal pleiotropy in causal relationships inferred from Mendelian randomization between complex traits and diseases. Nat Genet. 2018;50(5):693–8.
- Bowden J, Davey Smith G, Burgess S. Mendelian randomization with invalid instruments: effect estimation and bias detection through Egger regression. Int J Epidemiol. 2015;44(2):512–25.
- Sanderson E, Spiller W, Bowden J. Testing and correcting for weak and pleiotropic instruments in two-sample multivariable Mendelian randomization. Stat Med. 2021;40(25):5434–52.
- Zhang H, Ahearn TU, Lecarpentier J, Barnes D, Beesley J, Qi G, Jiang X, O'Mara TA, Zhao N, Bolla MK, et al. Genome-wide association study identifies 32 novel breast cancer susceptibility loci from overall and subtypespecific analyses. Nat Genet. 2020;52(6):572–81.
- Fernandez-Rozadilla C, Timofeeva M, Chen Z, Law P, Thomas M, Schmit S, Diez-Obrero V, Hsu L, Fernandez-Tajes J, Palles C, et al. Deciphering colorectal cancer genetics through multi-omic analysis of 100,204 cases and 154,587 controls of European and east Asian ancestries. Nat Genet. 2023;55(1):89–99.
- Scelo G, Purdue MP, Brown KM, Johansson M, Wang Z, Eckel-Passow JE, Ye Y, Hofmann JN, Choi J, Foll M, et al. Genome-wide association study identifies multiple risk loci for renal cell carcinoma. Nat Commun. 2017;8:15724.
- Phelan CM, Kuchenbaecker KB, Tyrer JP, Kar SP, Lawrenson K, Winham SJ, et al. Identification of 12 new susceptibility loci for different histotypes of epithelial ovarian cancer. Nat Genet. 2017;49(5):680–91.
- Long Y, Tang L, Zhou Y, Zhao S, Zhu H. Causal relationship between gut microbiota and cancers: a two-sample Mendelian randomisation study. BMC Med. 2023;21(1):1–14.
- Choi KW, Chen C-Y, Stein MB, Klimentidis YC, Wang M-J, Koenen KC, et al. Assessment of bidirectional relationships between physical activity and depression among adults: a 2-sample mendelian randomization study. JAMA Psychiat. 2019;76(4):399–408.
- Koppenol W, Bounds P, Dang C. Otto Warburg's contributions to current concepts of cancer metabolism, 117. Cuezva JM, Chen G, Alonso AM, et al., The bioenergetic signature of lung adenocarcinomas is a molecular marker of cancer diagnosis and prognosis. Carcinogenesis. 2004;25:1157–63.
- Pierce BL, Burgess S. Efficient design for Mendelian randomization studies: subsample and 2-sample instrumental variable estimators. Am J Epidemiol. 2013;178(7):1177–84.

- Hartwig FP, Davey Smith G, Bowden J. Robust inference in summary data Mendelian randomization via the zero modal pleiotropy assumption. Int J Epidemiol. 2017;46(6):1985–98.
- Li Y, Liu H, Ye S, Zhang B, Li X, Yuan J, Du Y, Wang J, Yang Y. The effects of coagulation factors on the risk of endometriosis: a Mendelian randomization study. BMC Med. 2023;21(1):195.
- Hemani G, Zheng J, Elsworth B, Wade KH, Haberland V, Baird D, Laurin C, Burgess S, Bowden J, Langdon R, et al. The MR-base platform supports systematic causal inference across the human phenome. Elife. 2018;7:e34408.
- Brion MJ, Shakhbazov K, Visscher PM. Calculating statistical power in Mendelian randomization studies. Int J Epidemiol. 2013;42(5):1497–501.
- Bulik-Sullivan B, Finucane HK, Anttila V, Gusev A, Day FR, Loh PR, ReproGen C, Psychiatric Genomics C, Genetic Consortium for Anorexia Nervosa of the Wellcome Trust Case Control C, Duncan L, et al. An atlas of genetic correlations across human diseases and traits. Nat Genet. 2015;47(11):1236–41.
- Zuber V, Grinberg NF, Gill D, Manipur I, Slob EA, Patel A, Wallace C, Burgess S: Combining evidence from Mendelian randomization and colocalization: review and comparison of approaches. Am J Human Genet 2022.
- Giambartolomei C, Vukcevic D, Schadt EE, Franke L, Hingorani AD, Wallace C, Plagnol V. Bayesian test for colocalisation between pairs of genetic association studies using summary statistics. PLoS Genet. 2014;10(5):e1004383.
- Chen J, Xu F, Ruan X, Sun J, Zhang Y, Zhang H, Zhao J, Zheng J, Larsson SC, Wang X, et al. Therapeutic targets for inflammatory bowel disease: proteome-wide Mendelian randomization and colocalization analyses. EBioMedicine. 2023;89:104494.
- Qiu S, Cai Y, Yao H, Lin C, Xie Y, Tang S, Zhang A. Small molecule metabolites: discovery of biomarkers and therapeutic targets. Signal Transduct Target Ther. 2023;8(1):132.
- 45. Stine ZE, Schug ZT, Salvino JM, Dang CV. Targeting cancer metabolism in the era of precision oncology. Nat Rev Drug Discov. 2022;21(2):141–62.
- Elia I, Haigis MC. Metabolites and the tumour microenvironment: from cellular mechanisms to systemic metabolism. Nat Metab. 2021;3(1):21–32.
- Pontremoli S, Melloni E, Viotti P, Michetti M, Di Lisa F, Siliprandi N. Isovalerylcarnitine is a specific activator of the high calcium requiring calpain forms. Biochem Biophys Res Commun. 1990;167(1):373–80.
- Suhre K, Dadhania DM, Lee JR, Muthukumar T, Chen Q, Gross SS, Suthanthiran M. Kidney allograft function is a confounder of urine metabolite profiles in kidney allograft recipients. Metabolites. 2021;11(8):533.
- Miranti EH, Freedman ND, Weinstein SJ, Abnet CC, Selhub J, Murphy G, Diaw L, Männistö S, Taylor PR, Albanes D. Prospective study of serum cysteine and cysteinylglycine and cancer of the head and neck, esophagus, and stomach in a cohort of male smokers. Am J Clin Nutr. 2016;104(3):686–93.
- Bade BC, Thomas DD, Scott JB, Silvestri GA. Increasing physical activity and exercise in lung cancer: reviewing safety, benefits, and application. J Thorac Oncol. 2015;10(6):861–71.
- Morris JS, Bradbury KE, Cross AJ, Gunter MJ, Murphy N. Physical activity, sedentary behaviour and colorectal cancer risk in the UK Biobank. Br J Cancer. 2018;118(6):920–9.
- Xu JY, Zhang C, Wang X, Zhai L, Ma Y, Mao Y, Qian K, Sun C, Liu Z, Jiang S, et al. Integrative proteomic characterization of human lung adenocarcinoma. Cell. 2020;182(1):245–61 (e217).
- Aune D, Chan DS, Lau R, Vieira R, Greenwood DC, Kampman E, Norat T. Dietary fibre, whole grains, and risk of colorectal cancer: systematic review and dose-response meta-analysis of prospective studies. BMJ. 2011;10:343.
- Lustgarten MS, Fielding RA. Metabolites associated with circulating interleukin-6 in older adults. J Gerontol A Biol Sci Med Sci. 2017;72(9):1277–83.

Publisher's Note

Springer Nature remains neutral with regard to jurisdictional claims in published maps and institutional affiliations.